

Expectations and Independence

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Theorem Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space, let $X_1, \dots, X_n: \Omega \rightarrow \mathbb{R}$ be random variables.

Then $X_1, \dots, X_n$ are independent $\iff$ for any Borel functions $f_1, \dots, f_n: \mathbb{R} \rightarrow \mathbb{R}$, $\mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] = \prod_{k=1}^n \mathbb{E}[f_k(X_k)]$

Recall lemma: **Lemma** Let $f: \Omega \rightarrow [0, \infty)$. Then there exist simple f_n 's ($f_n \geq 0$) s.t. $0 \leq f_n \uparrow f$, and s.t. f_n is $\sigma(f)/\mathcal{B}(\mathbb{R})$ -measurable for all $n \geq 1$.

Proof. Set $f_n = \sum_{k=0}^{2^n-1} \frac{k}{2^n} \cdot \mathbb{1}_{\{\frac{k}{2^n} \leq f < \frac{k+1}{2^n}\}}$. $\square$

Proof of Theorem First suppose that for any Borel functions $f_1, \dots, f_n: \mathbb{R} \rightarrow \mathbb{R}$, $\mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] = \prod_{k=1}^n \mathbb{E}[f_k(X_k)]$

Now fix sets $(E_k \in \sigma(X_k), 1 \leq k \leq n)$.

Since $\sigma(X_k) = \{X_k^{-1}(B) : B \in \mathcal{B}(\mathbb{R})\}$, we can find sets $B_1, \dots, B_n \in \mathcal{B}(\mathbb{R})$ such that $E_k = \{X_k \in B_k\}$ for $1 \leq k \leq n$.

Then taking $f_k = \mathbb{1}_{B_k}$, we have

$$\prod_{k=1}^n \mathbb{E}[f_k(X_k)] = \prod_{k=1}^n \mathbb{E}[\mathbb{1}_{B_k}(X_k)] = \prod_{k=1}^n \mathbb{P}(X_k \in B_k) = \prod_{k=1}^n \mathbb{P}(E_k)$$

$$\stackrel{||}{=} \mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] = \mathbb{E}\left[\prod_{k=1}^n \mathbb{1}_{B_k}(X_k)\right] = \mathbb{P}\left(\bigcap_{k=1}^n \{X_k \in B_k\}\right) = \mathbb{P}\left(\bigcap_{k=1}^n E_k\right). \text{ Thus } X_1, \dots, X_n \text{ are independent.}$$

Next suppose that $X_1, \dots, X_n$ are independent. Then for all $B_1, \dots, B_n \in \mathcal{B}_n$, $\mathbb{E}\left[\prod_{k=1}^n \mathbb{1}_{B_k}(X_k)\right] = \prod_{k=1}^n \mathbb{E}[\mathbb{1}_{B_k}(X_k)]$

Now let $f_1, \dots, f_n$ be simple Borel functions; we may write $f_k = \sum_{\ell=1}^m c_{k,\ell} \mathbb{1}_{B_{k,\ell}}$, for some $c_{k,\ell} \in \mathbb{R}$ and $B_{k,\ell} \in \mathcal{B}(\mathbb{R})$.

Then

(49)

$$\begin{aligned} \mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] &= \mathbb{E}\left[\prod_{k=1}^n \left(\sum_{l=1}^m c_{k,l} \mathbb{1}_{B_{k,l}}(X_k)\right)\right] \stackrel{\text{L.o.F.}}{=} \sum_{l_1, \dots, l_n=1}^m \mathbb{E}\left[\prod_{k=1}^n c_{k,l_k} \mathbb{1}_{B_{k,l_k}}(X_k)\right] \\ &= \sum_{l_1, \dots, l_n=1}^m \prod_{k=1}^n c_{k,l_k} \mathbb{E}\left[\mathbb{1}_{B_{k,l_k}}\right] = \prod_{k=1}^n \mathbb{E}\left[\sum_{l=1}^m c_{k,l} \mathbb{1}_{B_{k,l}}\right] = \prod_{k=1}^n \mathbb{E} f_k(X_k) \end{aligned}$$

So this gives the factorization formula for simple functions.

Next suppose $f_1, \dots, f_n$ are non-negative Borel functions, and set $Y_k = f_k(X_k)$. Then $Y_1, \dots, Y_n$ are independent. By the lemma, we may find simple functions $(Y_{n,m}, m \geq 1)$ s.t. $0 \leq Y_{n,m} \uparrow Y_n$ with $Y_{n,m}$ $\sigma(Y_n)/\mathcal{B}(\mathbb{R})$ -measurable.

By the factorization formula for simple functions, and the monotone convergence theorem,

$$\mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] = \mathbb{E}\left[\prod_{k=1}^n Y_k\right] \stackrel{\text{Mon. conv. thm.}}{=} \lim_{m \rightarrow \infty} \mathbb{E}\left[\prod_{k=1}^n Y_{k,m}\right] \stackrel{\text{F.F.}}{=} \lim_{m \rightarrow \infty} \prod_{k=1}^n \mathbb{E} Y_{k,m} \stackrel{\text{Mon. conv. thm.}}{=} \prod_{k=1}^n \mathbb{E} Y_k = \prod_{k=1}^n \mathbb{E} f_k(X_k).$$

This establishes the factorization formula for non-negative functions.

Finally, if $f_1, \dots, f_n$ are bounded and measurable then we can again use linearity to write

$f_k(X_k) = Y_k = Y_k^+ - Y_k^-$, and we then have

$$\begin{aligned} \mathbb{E}\left[\prod_{k=1}^n f_k(X_k)\right] &= \mathbb{E}\left[\prod_{k=1}^n (Y_k^+ - Y_k^-)\right] = \sum_{(\sigma_1, \dots, \sigma_n) \in \{+, -\}^n} (-1)^{\#\{i: \sigma_i = -\}} \mathbb{E}\left[\prod_{k=1}^n Y_k^{\sigma_k}\right] \\ &\stackrel{\text{non-neg. case}}{=} \sum_{(\sigma_1, \dots, \sigma_n) \in \{+, -\}^n} (-1)^{\#\{i: \sigma_i = -\}} \prod_{k=1}^n \mathbb{E}[Y_k^{\sigma_k}] = \prod_{k=1}^n \mathbb{E}[Y_k^+ - Y_k^-] = \prod_{k=1}^n \mathbb{E}[f_k(X_k)]. \quad \square \end{aligned}$$

Corollary If $X, Y: \Omega \rightarrow \mathbb{R}$ are independent and $\underbrace{X \in L^1(\mathbb{P}), Y \in L^1(\mathbb{P})}_{\mathbb{E}|X| < \infty, \mathbb{E}|Y| < \infty}$, then $XY \in L^1(\mathbb{P})$ and $\mathbb{E}XY = \mathbb{E}X \mathbb{E}Y$.

Proof Notation: for a function $f: \Omega \rightarrow \mathbb{R}$ write $f_{\leq n} := f \mathbb{1}_{|f| \leq n}$.

First suppose $X, Y \geq 0$. Then by Mon. conv. thm and independence,

$$\mathbb{E}[XY] = \lim_{n \rightarrow \infty} \mathbb{E}[X_{\leq n} Y_{\leq n}] = \lim_{n \rightarrow \infty} \mathbb{E}X_{\leq n} \mathbb{E}Y_{\leq n} = \mathbb{E}X \mathbb{E}Y.$$

Write $X = X^+ - X^-$, $Y = Y^+ - Y^-$.

Then $\mathbb{E}|XY| = \mathbb{E}X^+Y^+ + \mathbb{E}X^+Y^- + \mathbb{E}X^-Y^+ + \mathbb{E}X^-Y^-$
 non-neg. case $\rightarrow = (\mathbb{E}(X^+ + X^-))(\mathbb{E}(Y^+ + Y^-)) = \mathbb{E}|X| \mathbb{E}|Y| < \infty,$

So $XY \in L^1(\mathbb{P})$ and

$$\begin{aligned} \mathbb{E}XY &= \mathbb{E}X^+Y^+ - \mathbb{E}X^+Y^- - \mathbb{E}X^-Y^+ + \mathbb{E}X^-Y^- \\ &= (\mathbb{E}(X^+ - X^-))(\mathbb{E}(Y^+ - Y^-)) = \mathbb{E}X \mathbb{E}Y. \quad \square \end{aligned}$$

Monotone class theorem Fix a measurable space $(\Omega, \mathcal{F})$ let $\mathcal{P} \subset \mathcal{F}$ be a π -system over Ω , with $\Omega \in \mathcal{P}$. Let $\mathcal{A}$ be a collection of functions $f: \Omega \rightarrow \mathbb{R}$ st. the following all hold.

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- (a) $1_P \in \mathcal{A}$ for all $P \in \mathcal{P}$;
- (b) If $f, g \in \mathcal{A}$ and $c \in \mathbb{R}$ then $cf + g \in \mathcal{A}$
- (c) If $(f_n, n \geq 1)$ are in $\mathcal{A}$ and $0 \leq f_n \uparrow f$ and f is bounded then $f \in \mathcal{A}$.

Then $\mathcal{A} \supseteq \{f: \Omega \rightarrow \mathbb{R} : f \text{ is } (\sigma(\mathcal{P})/\mathcal{B}(\mathbb{R}))\text{-measurable}\}$.

Proof **Step 1** Let $\Lambda := \{F \subset \Omega : 1_F \in \mathcal{A}\}$. Then $\mathcal{P} \subset \Lambda$ by definition. ^(b)
 Also, if $E, F \in \Lambda$ and $E \subset F$ then $1_E, 1_F \in \mathcal{A}$ so $1_{F \setminus E} = 1_F - 1_E \in \mathcal{A}$ so $F \setminus E \in \Lambda$.
 Moreover, if $0 \leq E_n \uparrow E$ and each $E_n \in \Lambda$ then $1_{E_n} \uparrow 1_E$ so $1_E \in \mathcal{A}$ so $E \in \Lambda$.
 Thus Λ is a λ -system so $\sigma(\mathcal{P}) \subseteq \Lambda$ by Dynkin's π -system lemma.

Step 2 • By Step 1, For all $E \in \sigma(\mathcal{P})$, $1_E \in \mathcal{A}$.

- By (b), for any $E_1, \dots, E_n \in \sigma(\mathcal{P})$ and $c_1, \dots, c_n \in \mathbb{R}$, $\sum_{i=1}^n c_i 1_{E_i} \in \mathcal{A}$, so $\mathcal{A}$ contains all simple $(\sigma(\mathcal{P})/\mathcal{B}(\mathbb{R}))$ -measurable functions.
- If $f: \Omega \rightarrow [0, \infty)$ is $(\sigma(\mathcal{P})/\mathcal{B}(\mathbb{R}))$ -measurable then f is an increasing limit of simple f^n , $0 \leq f_n \uparrow f$, so if f is also bounded then $f \in \mathcal{A}$ by (c). So $\mathcal{A} \supseteq \{\text{Non-neg. bdd. } \sigma(\mathcal{P})/\mathcal{B}(\mathbb{R})\text{-meas. } f\}$
- Any $f: \Omega \rightarrow \mathbb{R}$ can be written as $f = f^+ - f^-$ and if f is $\sigma(\mathcal{P})/\mathcal{B}(\mathbb{R})$ meas. and bounded then f^+, f^- are $(\sigma(\mathcal{P})/\mathcal{B}(\mathbb{R}))$ -meas., non-neg. and bdd so $f^+, f^- \in \mathcal{A}$ so $f \in \mathcal{A}$ by (b). □

The probabilistic method

To prove an object with a given property P exists, one may construct a random variable X with the property that $P(X \text{ has property } P) > 0$.

Examples

① Existence of continuous nowhere differentiable functions

Let $D_n = \{\frac{i}{2^n}, 0 \leq i \leq 2^n\}$, $D_{\leq n} = \bigcup_{i \leq n} D_i$, $D = D_\infty$. Let $(N_x, x \in D)$ be i.i.d. $N(0,1)$ r.v.s.

Define $B_1(0) = 0$, $B_1(1) = N_1$, $B_n(x) = x N_1$ (linear interpolation)

Inductively, given B_{n-1} , let

$$B_n(x) = \begin{cases} B_{n-1}(x) & \text{if } x \in D_{\leq n-1} \\ B_{n-1}(x) + N_x / \sqrt{2^n} & \text{if } x \in D_n \setminus D_{\leq n-1} \\ \text{linear interpolation between points of } D_{\leq n}. \end{cases}$$

Then it turns out that $P(B_n \text{ is a uniformly convergent function}) = 1$.

The limit B_∞ is Brownian motion on $[0,1]$

And $P(B_\infty \text{ is nowhere differentiable}) = 1$.

② Small-discrepancy signings of vectors

Prop Let $v_1, \dots, v_n \in \mathbb{R}^n$ have $|v_i| = 1$. Then there exist $\sigma_1, \dots, \sigma_n \in \{-1, 1\}$ s.t. $|\sigma_1 v_1 + \dots + \sigma_n v_n| \leq \sqrt{n}$.

Proof

Let $\sigma_1, \dots, \sigma_n$ be independent and uniform on $\{-1, 1\}$. Set $X = |\sigma_1 v_1 + \dots + \sigma_n v_n|^2$.

Then $X = (\sigma_1 v_1 + \dots + \sigma_n v_n) \cdot (\sigma_1 v_1 + \dots + \sigma_n v_n)$

$$\begin{aligned} \mathbb{E}[X] &= \mathbb{E} \left[\sum_{i=1}^n \sum_{j=1}^n \sigma_i \sigma_j v_i \cdot v_j \right] = \sum_{i=1}^n \sum_{j=1}^n v_i \cdot v_j \mathbb{E}[\sigma_i \sigma_j] \\ &= \sum_{i=1}^n |v_i|^2 \mathbb{E}[\sigma_i^2] + 2 \sum_{1 \leq i < j \leq n} v_i \cdot v_j \mathbb{E}[\sigma_i \sigma_j] \\ &= \sum_{i=1}^n 1 + 2 \sum_{1 \leq i < j \leq n} 0 = n. \end{aligned}$$

So $\mathbb{E} X = n$ so $\mathbb{P}(|\sigma_1 v_1 + \dots + \sigma_n v_n| \leq \sqrt{n}) = \mathbb{P}(X \leq n) > 0$

So there must be at least one choice of $\sigma_1, \dots, \sigma_n$ s.t. $|\sigma_1 v_1 + \dots + \sigma_n v_n| \leq \sqrt{n}$.

... but how do we actually do stuff? $f_{\leq n} := f \mathbb{1}_{|f| \leq n}$

★ If X is a r.v. which takes values in $\mathbb{N}$ (or some other countable set) $N \subset [0, \infty)$ then $\mathbb{E}X = \lim_{n \rightarrow \infty} \mathbb{E}X_{\leq n} = \lim_{n \rightarrow \infty} \int X_{\leq n} d\mathbb{P}$; $X_{\leq n}$ is simple so

$$\mathbb{E}X_{\leq n} = \sum_{k=0}^{\infty} k \cdot \mathbb{P}(X=k) = \sum_{k=0}^{\infty} k \cdot \mu_X(k) \quad \text{by def. of } \mu_X \quad \text{and so } \mathbb{E}X = \sum_{k \in \mathbb{N}} k \cdot \mathbb{P}(X=k).$$

So, for example, if $P \sim \text{Poisson}(\lambda)$ then $\mathbb{E}P = \sum_{k \in \mathbb{N}} k \cdot \frac{\lambda^k e^{-\lambda}}{k!} = \lambda \cdot \sum_{k \geq 1} \frac{\lambda^{k-1} e^{-\lambda}}{(k-1)!} = \lambda$

★ Suppose I tell you $N = (N_i, i \geq 1)$ are IID $N(0,1)$ and $P \sim \text{Poisson}(1)$ is independent of N . How do you compute $\mathbb{P}\left(\sum_{i=1}^P N_i \geq 1\right) = \mathbb{E}\left[\mathbb{1}_{\sum_{i=1}^P N_i \geq 1}\right] = \int \mathbb{1}_{\sum_{i=1}^P N_i(\omega) \geq 1} P(d\omega)$?

Using densities, change of variables, and Fubini's Theorem.

Definitions Given a measure space $(\Omega, \mathcal{F}, \mu)$ and $f: \Omega \rightarrow \mathbb{R}$ non-negative and measurable, define a new measure μ^f on $(\Omega, \mathcal{F})$ by $\mu^f(A) = \int_A f d\mu := \int f \mathbb{1}_A d\mu$.

Examples

• Let X be a non-negative r.v. on $(\Omega, \mathcal{F}, \mathbb{P})$, for $\lambda \geq 0$ let $f_\lambda(x) = e^{-\lambda x}$

$$\text{Then } \mu_{f_\lambda}(X)(A) = \int_A e^{-\lambda x} d\mathbb{P} = \mathbb{E}[e^{-\lambda x} \mathbb{1}_A]$$

The function $\lambda \mapsto \mu_{f_\lambda}(X)(\Omega) = \mathbb{E}[e^{-\lambda x}]$ is the Laplace transform of X

Size-biasing

- Let X be a r.v. with $X \geq 0$, $\mathbb{E}X < \infty$. Recall μ_X is dist. of X .

Then $(\mathbb{R}, \mathcal{B}(\mathbb{R}), \mu_X)$ is a prob. space. The size-biased distribution of X

is $\hat{\mu}_X = (\mu_X \frac{X}{\mathbb{E}X})$, so $\hat{\mu}_X(A) = (\mu_X \frac{X}{\mathbb{E}X})(A) = \mathbb{E} \left[\frac{X}{\mathbb{E}X} \mathbb{1}_{X \in A} \right]$

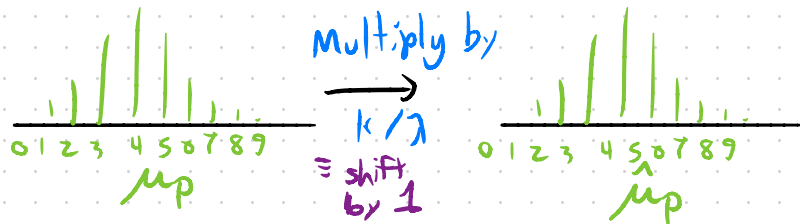
Note $\hat{\mu}_X(\mathbb{R}) = \mathbb{E} \left[\frac{X}{\mathbb{E}X} \right] = \frac{1}{\mathbb{E}X} \mathbb{E}[X] = 1$, so $\hat{\mu}_X$ is another prob. dist

- Let $P: \Omega \rightarrow \mathbb{R}$ be Poisson(λ); $\mathbb{P}(P=k) = \lambda^k e^{-\lambda} / k!$, $\mathbb{E}P = \lambda$.

Then $\hat{\mu}_P(A) = (\mu_P \frac{P}{\lambda})(A) = \mathbb{E} \left[\frac{P}{\lambda} \mathbb{1}_{P \in A} \right]$

$$= \sum_{k \in A \cap \{1, 2, 3, \dots\}} \frac{k}{\lambda} \cdot \mathbb{P}(P=k)$$

$$= \sum_{k \in A \cap \{1, 2, 3, \dots\}} \frac{k}{\lambda} \cdot \frac{\lambda^k e^{-\lambda}}{k!} = \sum_{k \in A \cap \{1, 2, 3, \dots\}} \frac{\lambda^{k-1} e^{-\lambda}}{(k-1)!} = \mathbb{P}(P+1 \in A) = \mu_{P+1}(A)$$



"A size-biased Poisson(λ) dist. is just $1 + \text{Poisson}(\lambda)$ "

Definition If $\nu = \mu f$ then we say ν has density f with respect to μ . (56)

Say X has density f with respect to Lebesgue measure if μ_X has density f w.r.t. Lebesgue, i.e. if $\forall B \in \mathcal{B}(\mathbb{R}), \mu_X(B) = \int \mathbb{1}_B \mu_X(dx) = \int_B f(x) dx$

The definition is justified by the following proposition.

Proposition If ν has density f with respect to μ then $\int g d\nu = \int g f d\mu$ holds for measurable g whenever a) $g \geq 0$ or b) $g \in L_1(\mu)$. Moreover, $g \in L^1(\nu) \Leftrightarrow g f \in L^1(\mu)$.

Proof If $g = \mathbb{1}_A$ for some $A \in \mathcal{F}$ then holds by definition. Then build up using monotone class thm and monotone convergence thm. $\square$

Change of variables formula

Let $(\Omega, \mathcal{F}, P)$ be a p.s. and let $X: \Omega \rightarrow \mathbb{R}$ be a r.v..

Then for all measurable $g: \mathbb{R} \rightarrow \mathbb{R}$ with $\mathbb{E}|g(X)| < \infty$, have $\mathbb{E} g(X) = \int_{\mathbb{R}} g(x) \mu(dx)$.

Moreover, if X has density f w.r.t. Lebesgue measure then $\mathbb{E} g(X) = \int_{\mathbb{R}} g(x) f(x) dx$.

Proof True for $\mathbb{1}_B(X)$, then build up using monotone class thm and monotone convergence thm. $\square$

Example If $N \sim N(0,1)$ then N has density $\Phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$ w.r.t. Lebesgue measure 57

$|N|$ has density $\Psi(x) = \sqrt{\frac{2}{\pi}} e^{-x^2/2} \mathbb{1}_{x \geq 0}$ w.r.t. Lebesgue measure.

What about the size-biasing of $|N|$?

The distribution is $\hat{\mu}_{|N|} = (\mu_{|N|} \cdot \frac{|N|}{\mathbb{E}|N|})$

$$\mathbb{E}|N| = \int |N| dP \underset{\text{c.f.v.}}{=} \int |x| d\mu_{|N|} \underset{\text{c.f.v.}}{=} \int_{[0,\infty)} x \cdot \sqrt{\frac{2}{\pi}} e^{-x^2/2} dx = \left[-\sqrt{\frac{2}{\pi}} e^{-x^2/2} \right]_0^\infty = \sqrt{\frac{2}{\pi}}.$$

So for B Borel, $B \subset [0,\infty)$,

$$\begin{aligned} \hat{\mu}_{|N|}(B) &= (\mu_{|N|} \cdot \frac{|N|}{\sqrt{2/\pi}})(B) = \int \mathbb{1}_B d(\mu_N \cdot \frac{|N|}{\mathbb{E}|N|}) = \int \mathbb{1}_B \cdot \frac{N}{\mathbb{E}|N|} d\mu \\ &= \int_B \frac{x}{\sqrt{2/\pi}} \sqrt{\frac{2}{\pi}} e^{-x^2/2} dx = \int_B x e^{-x^2/2} dx \end{aligned}$$

So $\hat{\mu}_{|N|}$ has density $x e^{-x^2/2} \mathbb{1}_{x \geq 0}$ w.r.t. Lebesgue measure.

This is the **Rayleigh distribution**.

Exercise If X, Y are independent standard Normal then $\sqrt{X^2 + Y^2}$ is Rayleigh distributed.

More generally, if $(\Omega, \mathcal{F}, \mu)$ measure space and $(S, \mathcal{A})$ is a measurable space, then for a $f: \Omega \rightarrow S$ an $(\mathcal{F}/\mathcal{A})$ -measurable function we may define $\nu(E) = \mu(f^{-1}(E))$ for $E \in \mathcal{A}$.

Then for all non-negative $(\mathcal{A}/\mathcal{B}(\mathbb{R}))$ -measurable functions $g: S \rightarrow \mathbb{R}$, $\int g d\nu = \int g \circ f d\mu$.

Special case: $\mathbb{E} g(X) = \int g(x) \mu(dx)$.

Another special case: if $X = (X_1, \dots, X_n)$ are real r.v.s defined on a common space and

$g: \mathbb{R}^n \rightarrow [0, \infty)$ is measurable then

$$\mathbb{E} g(X) = \int_{\mathbb{R}^n} g(\vec{x}) \mu_X(d\vec{x}) \quad \leftarrow \text{Makes sense since } X: \Omega \rightarrow \mathbb{R}^n \text{ is measurable so } \mu_X \text{ is defined and is a measure on } \mathbb{R}^n.$$

Exercise

Prove that if $\varphi: [a, b] \rightarrow \mathbb{R}$ is C_1 (continuously differentiable) and strictly $\uparrow$ then for all Borel functions $g: [\varphi(a), \varphi(b)] \rightarrow [0, \infty)$,

$$\int_{\varphi(a)}^{\varphi(b)} g(y) dy = \int_a^b g(\varphi(y)) \varphi'(y) dy.$$

(Note: If $g = \mathbb{1}_{[c, d]}$ for $[c, d] \subset [a, b]$ this says $\varphi(d) - \varphi(c) = \int_c^d \varphi'(y) dy$, which is the fundamental thm of calc. Then use monotone class theorem.)